

STAT 5010 Tutorial 2*

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Definitions

1. A parametric family $\{P_\theta : \theta \in \Theta\}$ dominated by a σ -finite measure ν on (Ω, \mathcal{F}) is called an *exponential family* if and only if

$$\frac{dP_\theta}{d\nu}(\omega) = \exp\{[\eta(\theta)]^\tau T(\omega) - \xi(\theta)\} h(\omega), \quad \omega \in \Omega,$$

where $\exp\{x\} = e^x$, T is a random p -vector with a fixed positive integer p , η is a function from Θ to \mathcal{R}^p , h is a nonnegative Borel function on (Ω, \mathcal{F}) , and $\xi(\theta) = \log \left\{ \int_\Omega \exp\{[\eta(\theta)]^\tau T(\omega)\} h(\omega) d\nu(\omega) \right\}$

2. In an exponential family, consider the reparameterization $\eta = \eta(\theta)$ and

$$f_\eta(\omega) = \exp\{\eta^\tau T(\omega) - \zeta(\eta)\} h(\omega), \quad \omega \in \Omega,$$

where $\zeta(\eta) = \log \left\{ \int_\Omega \exp\{\eta^\tau T(\omega)\} h(\omega) d\nu(\omega) \right\}$. This is the *canonical form* for the family, which is not unique. The new parameter η is called the *natural parameter*. The new parameter space $\Xi = \{\eta(\theta) : \theta \in \Theta\}$, a subset of \mathcal{R}^p , is called the *natural parameter space*. An exponential family in canonical form is called a *natural exponential family*. If there is an open set contained in the natural parameter space of an exponential family, then the family is said to be of *full rank*.

3. A measurable function of $X, T(X)$, is called a *statistic* if $T(X)$ is a known value whenever X is known, i.e., the function T is a known function.
4. Let X be a sample from an unknown population $P \in \mathcal{P}$, where \mathcal{P} is a family of populations. A statistic $T(X)$ is said to be *sufficient* for $P \in \mathcal{P}$ (or for $\theta \in \Theta$ when $\mathcal{P} = \{P_\theta : \theta \in \Theta\}$ is a parametric family) if and only if the conditional distribution of X given T is known (does not depend on P or θ).
5. Let T be a sufficient statistic for $P \in \mathcal{P}$. T is called a *minimal sufficient* statistic if and only if, for any other statistic S sufficient for $P \in \mathcal{P}$, there is a measurable function ψ such that $T = \psi(S)$ a.s. \mathcal{P}
6. A statistic $V(X)$ is said to be *ancillary* if its distribution does not depend on the population P and *first-order ancillary* if $E[V(X)]$ is independent of P .
7. A statistic $T(X)$ is said to be *complete* for $P \in \mathcal{P}$ if and only if, for any Borel f , $E[f(T)] = 0$ for all $P \in \mathcal{P}$ implies $f(T) = 0$ a.s. \mathcal{P} . T is said to be *boundedly complete* if and only if the previous statement holds for any bounded Borel f .

Propositions and Theorems

1. If η_0 is an interior point of the natural parameter space, then the m.g.f. ψ_{η_0} of $P_{\eta_0} \circ T^{-1}$ is finite in a neighborhood of 0 and is given by

$$\psi_{\eta_0}(t) = \exp\{\zeta(\eta_0 + t) - \zeta(\eta_0)\}$$

2. (The factorization theorem) Suppose that X is a sample from $P \in \mathcal{P}$ and \mathcal{P} is a family of probability measures on $(\mathcal{R}^n, \mathcal{B}^n)$ dominated by a σ -finite measure ν . Then $T(X)$ is sufficient for $P \in \mathcal{P}$ if and only if there are nonnegative Borel functions h (which does not depend on P) on $(\mathcal{R}^n, \mathcal{B}^n)$ and g_P (which depends on P) on the range of T such that

$$\frac{dP}{d\nu}(x) = g_P(T(x))h(x).$$

3. Suppose that \mathcal{P} contains p.d.f.'s f_P w.r.t. a σ -finite measure and that there exists a sufficient statistic $T(X)$ such that, for any possible values x and y of X , $f_P(x) = f_P(y)\phi(x, y)$ for all P implies $T(x) = T(y)$, where ϕ is a measurable function. Then $T(X)$ is minimal sufficient for $P \in \mathcal{P}$.
4. If P is in an exponential family of full rank with p.d.f.'s given as in Definition. 2, then $T(X)$ is complete and sufficient for $\eta \in \Xi$.

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5. (Basu's theorem) Let V and T be two statistics of X from a population $P \in \mathcal{P}$. If V is ancillary and T is boundedly complete and sufficient for $P \in \mathcal{P}$, then V and T are independent w.r.t. any $P \in \mathcal{P}$.
6. A complete and sufficient statistic is also minimal sufficient. However, a minimal sufficient statistic is not necessarily complete.

Question 1

1. Let X and Y be two random variables such that Y has the binomial distribution with size N and probability π and, given $Y = y$, X has the binomial distribution with size y and probability p . Suppose that $p \in (0, 1)$ and $\pi \in (0, 1)$ are unknown and N is known. Show that (X, Y) is minimal sufficient for (p, π) .

Solution: See the handout later.

Question 2

1. Let X_1, \dots, X_n be i.i.d. random variables from P_θ , the uniform distribution $U(\theta, \theta + 1)$, $\theta \in \mathcal{R}$. Prove that $T = (X_{(1)}, X_{(n)})$ is minimal sufficient.

Solution: See the handout later.

Proofs for Some propositions

- Prop. 3.

Proof. From Bahadur (1957), there exists a minimal sufficient statistic $S(X)$. The result follows if we can show that $T(X) = \psi(S(X))$ a.s. \mathcal{P} for a measurable function ψ . By the factorization theorem, there are Borel functions g_P and h such that $f_P(x) = g_P(S(x))h(x)$ for all P . Let $A = \{x : h(x) = 0\}$. Then $P(A) = 0$ for all P . For x and y such that $S(x) = S(y)$, $x \notin A$ and $y \notin A$,

$$\begin{aligned} f_P(x) &= g_P(S(x))h(x) \\ &= g_P(S(y))h(x)h(y)/h(y) \\ &= f_P(y)h(x)/h(y) \end{aligned}$$

for all P . Hence $T(x) = T(y)$. This shows that there is a function ψ such that $T(x) = \psi(S(x))$ except for $x \in A$. It remains to show that ψ is measurable. Since S is minimal sufficient, $g(T(X)) = S(X)$ a.s. \mathcal{P} for a measurable function g . Hence g is one-to-one and $\psi = g^{-1}$. The measurability of ψ follows from Theorem 3.9 in Parthasarathy (1967). ■

- Prop. 4.

Proof. Obviously, T is sufficient. Suppose that there is a function f such that $E[f(T)] = 0$ for all $\eta \in \Xi$. Then,

$$\int f(t) \exp\{\eta^\tau t - \zeta(\eta)\} d\lambda = 0 \quad \text{for all } \eta \in \Xi,$$

where λ is a measure on $(\mathcal{R}^p, \mathcal{B}^p)$. Let η_0 be an interior point of Ξ . Then

$$\int f_+(t) e^{\eta^\tau t} d\lambda = \int f_-(t) e^{\eta^\tau t} d\lambda \quad \text{for all } \eta \in N(\eta_0),$$

where $N(\eta_0) = \{\eta \in \mathcal{R}^p : \|\eta - \eta_0\| < \epsilon\}$ for some $\epsilon > 0$. In particular,

$$\int f_+(t) e^{\eta_0^\tau t} d\lambda = \int f_-(t) e^{\eta_0^\tau t} d\lambda = c.$$

If $c = 0$, then $f = 0$ a.e. λ . If $c > 0$, then $c^{-1}f_+(t)e^{\eta_0^\tau t}$ and $c^{-1}f_-(t)e^{\eta_0^\tau t}$ are p.d.f.'s w.r.t. λ and this implies that their m.g.f.'s are the same in a neighborhood of 0. Thus, $c^{-1}f_+(t)e^{\eta_0^\tau t} = c^{-1}f_-(t)e^{\eta_0^\tau t}$, i.e., $f = f_+ - f_- = 0$ a.e. λ . Hence T is complete. ■